

# Math 632 004 HW05

Shawn Zhong

TOTAL POINTS

**50 / 50**

## QUESTION 1

### 1 Durrett 2.4 10 / 10

✓ - 0 pts Correct

- 3 pts Minor errors. Examples: Incorrectly evaluated integrals, arithmetic or algebra mistakes
- 5 pts Substantial errors. Examples: integrals set up incorrectly
- 8 pts Serious errors. Examples: does not set up any appropriate integral, does not use exponential density at all
- 10 pts No work submitted

## QUESTION 2

### 2 Durrett 2.7abc 10 / 10

✓ - 0 pts Correct

- 1 pts (a) Minor error. Examples: arithmetic or algebra
- 3 pts (a) Little progress to solution. Examples: did not use properties of exponential distribution in any way
- 1 pts (b) Minor errors. Examples: computational mistakes
- 2 pts (b) Substantial errors. Examples: integral not set up correctly
- 4 pts (b) Little progress to solution. Examples: did not use properties of exponential distribution in any way
- 1 pts (c) Minor errors. Examples: computational mistakes
- 3 pts (c) Little progress to solution. Examples: did not use properties of exponential distribution in any way
- 10 pts No work submitted

## QUESTION 3

### 3 Durrett 2.9b 10 / 10

✓ - 0 pts Correct

- 3 pts Minor errors. Examples: arithmetic or algebra errors
- 5 pts Substantial errors. Examples: incorrect expression for Bob's total waiting time
- 8 pts Little progress is made towards the solution
- 10 pts No work submitted

## QUESTION 4

### 4 Durrett 2.13 10 / 10

✓ - 0 pts Correct

- 1 pts (a) Minor computational error
- 2 pts (a) Substantial error. Examples: incorrectly set up probability in terms of the Poisson processes
- 4 pts (a) Little progress to the solution
- 1 pts (b) Minor computational error
- 3 pts (b) Substantial error
- 1 pts (c) Incorrect conclusion due to minor error
- 3 pts (c) Serious errors: Incorrect definition of independence
- 10 pts No work submitted

## QUESTION 5

### 5 Problem 5 10 / 10

✓ - 0 pts Correct

- 3 pts Computational errors
- 5 pts Substantial error. Examples: Mistake in integral set up
- 8 pts Little progress to the solution
- 10 pts No work submitted
- 3 pts Forgot the summation

# HW5 - Problem & Solution

Thursday, October 25, 2018 9:47 PM

Math 632 Lecture 004 (Shinault), Fall 2018, Homework 5

Due Thursday November 1 by 10:05am

## Question 1

**2.4.** Copy machine 1 is in use now. Machine 2 will be turned on at time  $t$ . Suppose that the machines fail at rate  $\lambda_i$ . What is the probability that machine 2 is the first to fail?

- Let  $T_1 \sim \text{Exp}(\lambda_1)$  and  $T_2 \sim \text{Exp}(\lambda_2)$  be the running time for two machines until they fail
- $\mathbb{P}(T_2 + t < T_1) = \int_0^\infty f_{T_2}(s) \mathbb{P}(T_1 > s + t) ds = \int_0^\infty \lambda_2 e^{-\lambda_2 s} \cdot e^{-\lambda_1(s+t)} ds = \frac{\lambda_2}{\lambda_1 + \lambda_2} e^{-\lambda_1 t}$

## Question 2

Durrett Exercise 2.7 (a), (b), (c), page 93. (Ignore part (d).) In part (b), compute first  $P(V - U > s)$  for  $s > 0$  either by integrating densities of  $S$  and  $T$  or by conditioning on the events  $S < T$  and  $T < S$ . From  $P(V - U > s)$  deduce the density function  $f_{V-U}$  of  $V - U$ , and then the mean  $E(V - U)$  by integrating the density. Finally, check that your answers to (a), (b), (c) satisfy  $E(V - U) = E(V) - E(U)$ .

**2.7.** Let  $S$  and  $T$  be exponentially distributed with rates  $\lambda$  and  $\mu$ . Let  $U = \min\{S, T\}$  and  $V = \max\{S, T\}$ . Find (a)  $EU$ . (b)  $E(V - U)$ , (c)  $EV$ . (d) Use the identity  $V = S + T - U$  to get a different looking formula for  $EV$  and verify the two are equal.

### Part (a)

- $\mathbb{P}(U > x) = \mathbb{P}(S, T > x) = \mathbb{P}(S > x) \mathbb{P}(T > x) = e^{-(\lambda+\mu)x} \Rightarrow U \sim \text{Exp}(\lambda + \mu) \Rightarrow \mathbb{E}[U] = \frac{1}{\lambda + \mu}$

### Part (b)

- Claim:  $\mathbb{P}(X - Y > s) = \frac{\lambda_2}{\lambda_1 + \lambda_2} e^{-\lambda_1 s}$  for  $X \sim \text{Exp}(\lambda_1), Y \sim \text{Exp}(\lambda_2)$  and  $s > 0$ 
  - $\mathbb{P}(X > Y + s) = \int_0^\infty f_Y(y) \mathbb{P}(X > y + s) dy = \int_0^\infty \lambda_2 e^{-\lambda_2 y} \cdot e^{-\lambda_1(y+s)} dy = \frac{\lambda_2}{\lambda_1 + \lambda_2} e^{-\lambda_1 s}$
- $\mathbb{P}(V - U > x) = \mathbb{P}(T - S > x) + \mathbb{P}(S - T > x) = \frac{\lambda}{\lambda + \mu} e^{-\mu x} + \frac{\mu}{\lambda + \mu} e^{-\lambda x}$ , for  $x > 0$
- $f_{V-U}(x) = -\frac{d}{dx} \left( \frac{\lambda}{\lambda + \mu} e^{-\mu x} + \frac{\mu}{\lambda + \mu} e^{-\lambda x} \right) = \frac{\lambda\mu}{\lambda + \mu} (e^{-\mu x} + e^{-\lambda x})$ , for  $x > 0$
- $\mathbb{E}[V - U] = \int_0^\infty x \cdot f_{V-U}(x) dx = \int_0^\infty x \frac{\lambda\mu}{\lambda + \mu} (e^{-\mu x} + e^{-\lambda x}) dx = \frac{\lambda\mu}{\lambda + \mu} \left( \frac{1}{\mu^2} + \frac{1}{\lambda^2} \right)$

### Part (c)

- $\mathbb{P}(V < x) = \mathbb{P}(S, T < x) = \mathbb{P}(S < x) \mathbb{P}(T < x) = (1 - e^{-x\lambda})(1 - e^{-x\mu})$

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✓ - **0 pts** Correct

- **3 pts** Minor errors. Examples: Incorrectly evaluated integrals, arithmetic or algebra mistakes

- **5 pts** Substantial errors. Examples: integrals set up incorrectly

- **8 pts** Serious errors. Examples: does not set up any appropriate integral, does not use exponential density at all

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## Question 2

Durrett Exercise 2.7 (a), (b), (c), page 93. (Ignore part (d).) In part (b), compute first  $P(V - U > s)$  for  $s > 0$  either by integrating densities of  $S$  and  $T$  or by conditioning on the events  $S < T$  and  $T < S$ . From  $P(V - U > s)$  deduce the density function  $f_{V-U}$  of  $V - U$ , and then the mean  $E(V - U)$  by integrating the density. Finally, check that your answers to (a), (b), (c) satisfy  $E(V - U) = E(V) - E(U)$ .

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### Part (a)

- $\mathbb{P}(U > x) = \mathbb{P}(S, T > x) = \mathbb{P}(S > x) \mathbb{P}(T > x) = e^{-(\lambda+\mu)x} \Rightarrow U \sim \text{Exp}(\lambda + \mu) \Rightarrow \mathbb{E}[U] = \frac{1}{\lambda + \mu}$

### Part (b)

- Claim:  $\mathbb{P}(X - Y > s) = \frac{\lambda_2}{\lambda_1 + \lambda_2} e^{-\lambda_1 s}$  for  $X \sim \text{Exp}(\lambda_1), Y \sim \text{Exp}(\lambda_2)$  and  $s > 0$ 
  - $\mathbb{P}(X > Y + s) = \int_0^\infty f_Y(y) \mathbb{P}(X > y + s) dy = \int_0^\infty \lambda_2 e^{-\lambda_2 y} \cdot e^{-\lambda_1(y+s)} dy = \frac{\lambda_2}{\lambda_1 + \lambda_2} e^{-\lambda_1 s}$
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- $f_{V-U}(x) = -\frac{d}{dx} \left( \frac{\lambda}{\lambda + \mu} e^{-\mu x} + \frac{\mu}{\lambda + \mu} e^{-\lambda x} \right) = \frac{\lambda\mu}{\lambda + \mu} (e^{-\mu x} + e^{-\lambda x})$ , for  $x > 0$
- $\mathbb{E}[V - U] = \int_0^\infty x \cdot f_{V-U}(x) dx = \int_0^\infty x \frac{\lambda\mu}{\lambda + \mu} (e^{-\mu x} + e^{-\lambda x}) dx = \frac{\lambda\mu}{\lambda + \mu} \left( \frac{1}{\mu^2} + \frac{1}{\lambda^2} \right)$

### Part (c)

- $\mathbb{P}(V < x) = \mathbb{P}(S, T < x) = \mathbb{P}(S < x) \mathbb{P}(T < x) = (1 - e^{-x\lambda})(1 - e^{-x\mu})$

- $f_V(x) = \frac{d}{dx}[(1 - e^{-x\lambda})(1 - e^{-x\mu})] = \lambda e^{-x\lambda} + \mu e^{-x\mu} - (\lambda + \mu)e^{-x(\lambda+\mu)}$
- $\mathbb{E}[V] = \int_0^\infty x \cdot f_V(x) dx = \int_0^\infty x(\lambda e^{-x\lambda} + \mu e^{-x\mu} - (\lambda + \mu)e^{-x(\lambda+\mu)}) dx = \frac{1}{\lambda} + \frac{1}{\mu} - \frac{1}{\lambda + \mu}$
- Verify:  $\mathbb{E}[V] - \mathbb{E}[U] = \left(\frac{1}{\lambda} + \frac{1}{\mu} - \frac{1}{\lambda + \mu}\right) - \frac{1}{\lambda + \mu} = \frac{\lambda_2 + \mu^2}{\lambda\mu(\lambda + \mu)} = \frac{\lambda\mu}{\lambda + \mu} \left(\frac{1}{\mu^2} + \frac{1}{\lambda^2}\right) = \mathbb{E}[V - U]$

### Question 3

Durrett Exercise 2.9 (b), page 93. (Ignore part (a).) For systematic notation, let  $A_i$  denote the amount of time Al spends at server  $i$ , and  $B_i$  the amount of time Bob spends at server  $i$ , for  $i = 1, 2$ .

*Hint.* Draw a picture of the time line to understand how Al and Bob move through the servers. Use your answer from 2.7. This problem requires no integration.

**2.9.** In a hardware store you must first go to server 1 to get your goods and then go to a server 2 to pay for them. Suppose that the times for the two activities are exponentially distributed with means 6 minutes and 3 minutes. (a) Compute the average amount of time it take Bob to get his goods and pay if when he comes in there is one customer named Al with server 1 and no one at server 2. (b) Find the answer when times for the two activities are exponentially distributed with rates  $\lambda$  and  $\mu$ .

- Let  $A_1, B_1 \sim \text{Exp}(\lambda)$  and  $A_2, B_2 \sim \text{Exp}(\mu)$  be mutually independent
- If  $B_1 \leq A_2$ , then the total waiting time is  $A_1 + A_2 + B_2$
- If  $B_1 > A_2$ , then the total waiting time is  $A_1 + B_1 + B_2$
- Therefore,  $\mathbb{E}[A_1 + \max\{B_1, A_2\} + B_2] = \frac{1}{\lambda} + \left(\frac{1}{\lambda} + \frac{1}{\mu} - \frac{1}{\lambda + \mu}\right) + \frac{1}{\mu} = \frac{2}{\lambda} + \frac{2}{\mu} - \frac{1}{\lambda + \mu}$

### Question 4

Durrett Exercise 2.13, page 93.

*Hint.* Let  $N_i$  denote a Poisson process of rate  $\lambda_i$ . For part (a), express the event in terms of  $N_1, N_2$  and  $N_3$ . You may assume that the shocks come independently, so  $N_1, N_2$  and  $N_3$  are independent. Part (b) follows quickly from (a). For part (c) check whether  $P(U > s, V > t)$  equals  $P(U > s) \cdot P(V > t)$ .

**2.13.** A machine has two critically important parts and is subject to three different types of shocks. Shocks of type  $i$  occur at times of a Poisson process with rate  $\lambda_i$ . Shocks of types 1 break part 1, those of type 2 break part 2, while those of type 3 break both parts. Let  $U$  and  $V$  be the failure times of the two parts. (a) Find  $P(U > s, V > t)$ . (b) Find the distribution of  $U$  and the distribution of  $V$ . (c) Are  $U$  and  $V$  independent?

#### Part (a)

- $\mathbb{P}(U > s, V > t) = \mathbb{P}(N_1(s) = 0, N_2(t) = 0, N_3(\max\{s, t\}) = 0) = e^{-\lambda_1 s} e^{-\lambda_2 t} e^{-\lambda_3 \max\{s, t\}}$

#### Part (b)

- $\mathbb{P}(U > s) = \mathbb{P}(N_1(s) = 0, N_3(s) = 0) = e^{-\lambda_1 s} e^{-\lambda_3 s} = e^{-(\lambda_1 + \lambda_3)s} \Rightarrow f_U(s) = 1 - e^{-(\lambda_1 + \lambda_3)s}$
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## 2 Durrett 2.7abc 10 / 10

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- 1 pts (a) Minor error. Examples: arithmetic or algebra
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- $f_V(x) = \frac{d}{dx}[(1 - e^{-x\lambda})(1 - e^{-x\mu})] = \lambda e^{-x\lambda} + \mu e^{-x\mu} - (\lambda + \mu)e^{-x(\lambda+\mu)}$
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### Question 3

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### Question 4

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*Hint.* Let  $N_i$  denote a Poisson process of rate  $\lambda_i$ . For part (a), express the event in terms of  $N_1, N_2$  and  $N_3$ . You may assume that the shocks come independently, so  $N_1, N_2$  and  $N_3$  are independent. Part (b) follows quickly from (a). For part (c) check whether  $P(U > s, V > t)$  equals  $P(U > s) \cdot P(V > t)$ .

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#### Part (a)

- $\mathbb{P}(U > s, V > t) = \mathbb{P}(N_1(s) = 0, N_2(t) = 0, N_3(\max\{s, t\}) = 0) = e^{-\lambda_1 s} e^{-\lambda_2 t} e^{-\lambda_3 \max\{s, t\}}$

#### Part (b)

- $\mathbb{P}(U > s) = \mathbb{P}(N_1(s) = 0, N_3(s) = 0) = e^{-\lambda_1 s} e^{-\lambda_3 s} = e^{-(\lambda_1 + \lambda_3)s} \Rightarrow f_U(s) = 1 - e^{-(\lambda_1 + \lambda_3)s}$
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### Question 3

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- $\mathbb{P}(V > t) = \mathbb{P}(N_2(t) = 0, N_3(t) = 0) = e^{-\lambda_2 t} e^{-\lambda_3 t} = e^{-(\lambda_2 + \lambda_3)t} \Rightarrow f_V(t) = 1 - e^{-(\lambda_2 + \lambda_3)t}$

### Part (c)

- $\mathbb{P}(U > s)\mathbb{P}(V > t) = e^{-(\lambda_1+\lambda_3)s}e^{-(\lambda_2+\lambda_3)t} \neq e^{-\lambda_1 s}e^{-\lambda_2 t}e^{-\lambda_3 \max\{s,t\}} = \mathbb{P}(U > s, V > t)$
- Therefore  $U$  and  $V$  are not independent

### Question 5

Consider a Poisson process of rate  $\lambda$  and let  $s$  be a fixed positive number. Let  $\sigma$  be the random amount of time from  $s$  till the next arrival. In symbols,

$$\sigma = \left( \min_{k: T_k > s} T_k \right) - s.$$

Calculate rigorously the probability  $P(\sigma > a)$  for real  $a > 0$  by using the density functions of the arrival times  $T_k$  and the interarrival times  $\tau_k$ .

*Hint.* Draw a picture of the time line. and express the event  $\{\sigma > a\}$  in terms of the arrival times  $\{T_k\}$  whose density functions we know. Since  $s + \sigma$  is one of the arrival times, you can decompose the probability  $P(\sigma > a)$  into different cases according to which  $T_k$  is equal to  $s + \sigma$ .

$$\begin{aligned} \bullet \quad \mathbb{P}(\sigma > a) &= \mathbb{P}(s + \sigma > s + a) = \sum_{k=0}^{\infty} \mathbb{P}(T_k < s < a + s < T_{k+1}) \\ &= \sum_{k=0}^{\infty} \mathbb{P}(\tau_{k+1} > a + s - T_k, T_k < s) = \sum_{k=0}^{\infty} \left( \int_0^s f_{T_k}(t) \mathbb{P}(\tau_{k+1} > a + s - t) dt \right) \\ &= \sum_{k=0}^{\infty} \left( \int_0^s \left( \lambda e^{-\lambda t} \frac{(\lambda t)^{k-1}}{(k-1)!} \right) (e^{-\lambda(a+s-t)}) dt \right) = \sum_{k=0}^{\infty} \left( e^{-(a+s)\lambda} \lambda^k \int_0^s \frac{t^{k-1}}{(k-1)!} dt \right) \\ &= \sum_{k=0}^{\infty} \left( e^{-(a+s)\lambda} \lambda^k \left[ \frac{t^k}{k!} \right]_{t=0}^{t=s} \right) = e^{-\lambda(a+s)} \sum_{k=0}^{\infty} \frac{(\lambda s)^k}{k!} = e^{-\lambda(a+s)} \cdot e^{\lambda s} = e^{-\lambda a} \end{aligned}$$

#### 4 Durrett 2.13 10 / 10

✓ - **0 pts** Correct

- **1 pts** (a) Minor computational error
- **2 pts** (a) Substantial error. Examples: incorrectly set up probability in terms of the Poisson processes
- **4 pts** (a) Little progress to the solution
- **1 pts** (b) Minor computational error
- **3 pts** (b) Substantial error
- **1 pts** (c) Incorrect conclusion due to minor error
- **3 pts** (c) Serious errors: Incorrect definition of independence
- **10 pts** No work submitted

### Part (c)

- $\mathbb{P}(U > s)\mathbb{P}(V > t) = e^{-(\lambda_1+\lambda_3)s}e^{-(\lambda_2+\lambda_3)t} \neq e^{-\lambda_1s}e^{-\lambda_2t}e^{-\lambda_3 \max\{s,t\}} = \mathbb{P}(U > s, V > t)$
- Therefore  $U$  and  $V$  are not independent

### Question 5

Consider a Poisson process of rate  $\lambda$  and let  $s$  be a fixed positive number. Let  $\sigma$  be the random amount of time from  $s$  till the next arrival. In symbols,

$$\sigma = \left( \min_{k: T_k > s} T_k \right) - s.$$

Calculate rigorously the probability  $P(\sigma > a)$  for real  $a > 0$  by using the density functions of the arrival times  $T_k$  and the interarrival times  $\tau_k$ .

*Hint.* Draw a picture of the time line. and express the event  $\{\sigma > a\}$  in terms of the arrival times  $\{T_k\}$  whose density functions we know. Since  $s + \sigma$  is one of the arrival times, you can decompose the probability  $P(\sigma > a)$  into different cases according to which  $T_k$  is equal to  $s + \sigma$ .

$$\begin{aligned} \bullet \quad \mathbb{P}(\sigma > a) &= \mathbb{P}(s + \sigma > s + a) = \sum_{k=0}^{\infty} \mathbb{P}(T_k < s < a + s < T_{k+1}) \\ &= \sum_{k=0}^{\infty} \mathbb{P}(\tau_{k+1} > a + s - T_k, T_k < s) = \sum_{k=0}^{\infty} \left( \int_0^s f_{T_k}(t) \mathbb{P}(\tau_{k+1} > a + s - t) dt \right) \\ &= \sum_{k=0}^{\infty} \left( \int_0^s \left( \lambda e^{-\lambda t} \frac{(\lambda t)^{k-1}}{(k-1)!} \right) (e^{-\lambda(a+s-t)}) dt \right) = \sum_{k=0}^{\infty} \left( e^{-(a+s)\lambda} \lambda^k \int_0^s \frac{t^{k-1}}{(k-1)!} dt \right) \\ &= \sum_{k=0}^{\infty} \left( e^{-(a+s)\lambda} \lambda^k \left[ \frac{t^k}{k!} \right]_{t=0}^{t=s} \right) = e^{-\lambda(a+s)} \sum_{k=0}^{\infty} \frac{(\lambda s)^k}{k!} = e^{-\lambda(a+s)} \cdot e^{\lambda s} = e^{-\lambda a} \end{aligned}$$

## 5 Problem 5 10 / 10

✓ - 0 pts Correct

- 3 pts Computational errors
- 5 pts Substantial error. Examples: Mistake in integral set up
- 8 pts Little progress to the solution
- 10 pts No work submitted
- 3 pts Forgot the summation